Learning and Inference for Graphical and Hierarchical Models: A Personal Journey

> Alan S. Willsky willsky@mit.edu http://lids.mit.edu http://ssg.mit.edu

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Undirected graphical models

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•
$$\boldsymbol{G} = (V, E) (V = vertices; E \subset V \times V = edges)_{1}$$

- Markovianity on *G*
 - Hammersley-Clifford (NASC for positive dist.)
 - $\mathbf{C} = \text{Set of all cliques in } \boldsymbol{G}$

$$p(x) = \frac{1}{Z} \exp \left\{ \sum_{\mathcal{C} \in \mathbf{C}} \varphi_{\mathcal{C}}(x_{\mathcal{C}}) \right\}$$

• φ_c = Clique potential Z = partition function

Pairwise graphical models

$$p(x) = \frac{1}{Z} \exp\left\{\sum_{s \in V} \varphi_s(x_s) + \sum_{(s,t) \in E} \varphi_{st}(x_s, x_t)\right\}$$
$$p(x) = \frac{1}{Z} \prod_{s \in V} \psi_s(x_s) \prod_{(s,t) \in E} \psi_{st}(x_s, x_t)$$

Directing/undirecting the graphs?

- Undirected models: Factors are *not* typically probabilities
 Although they *can* be for cycle-free graphs (see BP), i.e., *trees*
- Directed models: Specify in terms of transition probabilities (parents to children)
- Directed to undirected: Easy (after moralization)
- Undirected to directed: Hard (and often a mistake) unless the graph is a tree (see BP)

Gaussian models

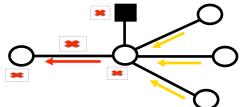
- $X \sim N(\mu, P)$ or $N^{-1}(h, J)$, with $J = P^{-1}$ and $h = J\mu$
- The sparsity structure of J determines graph structure
 - I.e., $J_{st} = 0$ if (s, t) is not an edge
- Directed model (0-mean for simplicity):

$$AX = W$$

- *W* ~ N(0, *I*)
- A Lower triangular
- $A = J^{1/2} \rightarrow$ In general we get lots of *fill*, unless the graph structure is a tree
- And it's even more complicated for non-Gaussian models, as higher-order cliques are introduced

Belief Propagation: Message passing for pairwise models on *trees*

Fixed point equations for likelihoods from disjoint parts of the graph:



$$m_{ts}(x_s) = \alpha \int_{x_t} \psi_{s,t}(x_s, x_t) \psi_t(x_t) \prod_{u \in \mathcal{N}(t) \setminus s} m_{ut}(x_t) dx_t$$
$$P_s(x_s) = \alpha \psi_s(x_s) \prod_{u \in \mathcal{N}(s)} m_{us}(x_s)$$

$$P_{st}(x_s, x_t) = \alpha \psi_{st}(x_s, x_t) \psi_s(x_s) \psi_t(x_t)$$
$$\prod_{u \in \mathcal{N}(s)/t} m_{us}(x_s) \prod_{u \in \mathcal{N}(t)/s} m_{ut}(x_t)$$

BP on trees

Gives factored form for distribution in terms of probability distributions

$$P(\lbrace x_s | s \in V \rbrace) = \prod_{s \in V} P_s(x_s) \prod_{(s,t) \in E} \frac{P_{st}(x_s, x_t)}{P_s(x_s) P_t(x_t)}$$

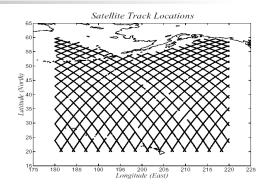
- Great flexibility in message-scheduling
 - Leaves-root-leaves = Rauch-Tung-Striebel
 - Completely parallel messaging, convergence in number of steps = diameter of the graph

Modeling of *structure*: Four questions

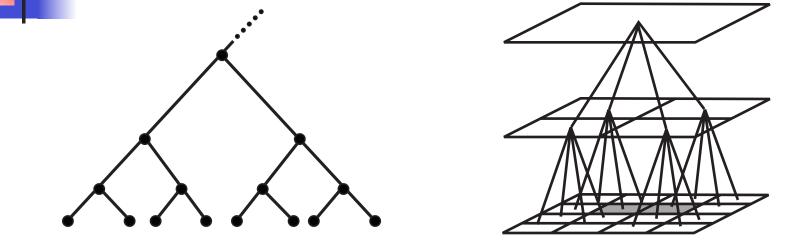
- What should the structure of the graph be?
- Which variables go where?
- What about adding hidden (unobserved) nodes (and variables)?
- What should the dimensions of the various variables be?

Our initial foray (with thanks to Michèle Basseville and Albert Benveniste): Multiresolution Models

- What can be multiresolution?
 - The *phenomenon* being modeled
 - The data that are collected
 - The *objectives* of statistical inference
 - The algorithms that result
- Some applications that motivated us (and others)
 - Oceanography
 - Groundwater hydrology
 - "Fractal priors" in regularization formulations in computer vision, mathematical physics, ...
 - Texture discrimination
 - Helioseismology (?) ...



Specifying MR models on trees



- MR synthesis, leads, as with Markov chains, to thinking about directed trees:
- E.g.: x(s) = A(s)x(sy) + w(s)
 - E.g.: Midpoint deflection is such a model
- Note that the *dimension* of the variables comes into play
- But let's assume we pre-specify the tree structure

A control theorist's idea: Internal models

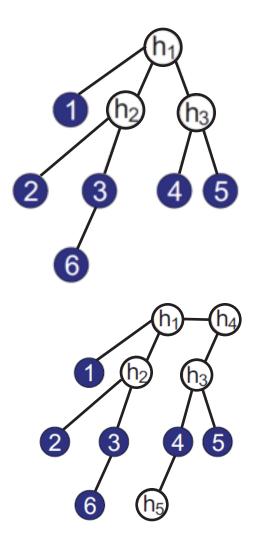
- Variables at coarser nodes are *linear functionals* of the finest-scale variables
 - Some of these may be measured or important quantities to be estimated
 - The rest are to be *designed*
 - To approximate the condition for tree-Markovianity
 - To yield a model that is "close" to the true fine-scale statistics
 - Scale-recursive algebraic design
 - Criterion used: Canonical correlations or predictive efficiency
 - Alternate using midpoint deflection or wavelets**
- Confounding the control theorist
 - Internal models need not have minimal state dimension

$$x(s) = x(s\overline{\gamma}) + w(s)$$

So, what if we want a tree but not necessarily a hierarchical one

- One approach: Construct the maximum likelihood tree given sample data (or the full second-order statistics)
- NOTE: This is quite different from what system theorists typically consider
 - There is no "state" to identify: All of the variables in the model we desire are observed
 - It is the *index set* that we get to play with
- Chow-Liu found a very efficient algorithm
 - Form graph with each observed variable at a different node
 - Edge weight between any two variables is their mutual information
 - Compute max-weight spanning tree
- What if we want hidden/latent "states"?

Reconstruction of a Latent Tree

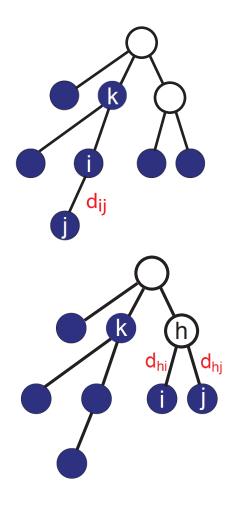


- Reconstruct a latent tree using exact statistics (first) or samples (to follow) of observed nodes only
- Exact/consistent recovery of *minimal* latent trees
 - Each hidden node has at least 3 neighbors
 - Observed variables are neither perfectly dependent nor independent
- Other objectives:
 - Computational efficiency
 - Low sample complexity

Information Distance

- Gaussian distributions $d_{ij} := -\log |\rho_{ij}|$ $\rho_{ij} := \frac{\operatorname{Cov}(X_i, X_j)}{\sqrt{\operatorname{Var}(X_i)\operatorname{Var}(X_j)}}$
- Discrete distributions $d_{ij} := -\log \frac{|\det \mathbf{J}^{ij}|}{\sqrt{\det \mathbf{M}^i \det \mathbf{M}^j}}$ $\mathbf{J}^{ij} \text{ Joint probability matrix} \qquad \mathbf{M}^i \text{ Marginal probability matrix (diagonal)}$ $\mathbf{Additivity}$ $d_{k,l} = \sum_{(i,j)\in \text{Path}((k,l);E_p)} d_{i,j} \qquad \mathbf{J}^{ij} = d_{1h_1} + d_{h_1h_2} + d_{2h_2}$

Testing Node Relationships



Node j – a leaf node Node i – parent of j for all $k \neq i, j$ $d_{jk} - d_{ik} = d_{ij}$ Can identify (parent, leaf child) pair

Node i and j – leaf nodes and share the same parent (sibling nodes) \iff for all $k \neq i, j \quad d_{jk} - d_{ik} = d_{hj} - d_{hi}$

Can identify leaf-sibling pairs.

Recursive Grouping (exact statistics)

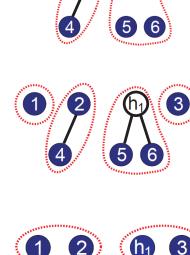
Step 1. Compute $d_{jk} - d_{ik}$ for all observed nodes (i, j, k).

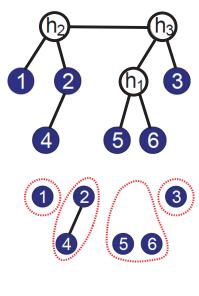
Step 2. Identify (parent, leaf child) or (leaf siblings) pairs.

Step 3. Introduce a hidden parent node for each sibling group without a parent.

Step 4. Compute the information distance for new hidden nodes. E.g.: $d_{5h_1} = \frac{1}{2}(d_{56} + d_{53} - d_{63})$

Step 5. Remove the identified child nodes and repeat Steps 2-4.

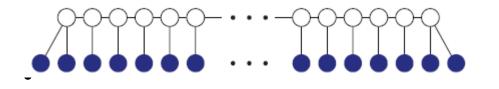




Recursive Grouping

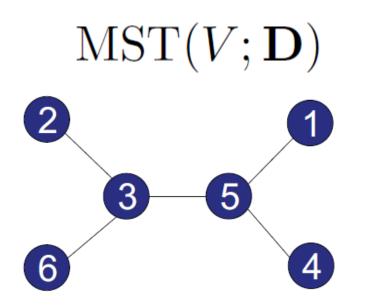
- Identifies a group of family nodes at each step.
- Introduces hidden nodes recursively.
- Correctly recovers all minimal latent trees.

• Computational complexity O(diam(T) m³).



Worst case O(m⁴)

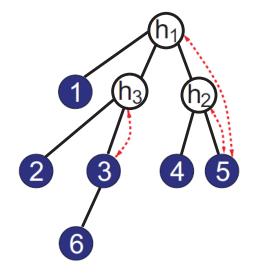
Chow-Liu Tree



Minimum spanning tree of V using D as edge weights

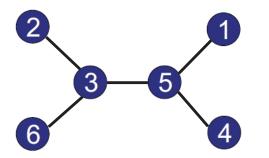
- V = set of observed nodes
- D = information distances
- Computational complexity O(m² log m)

Surrogate Nodes and the C-L Tree



V = set of observed nodes Surrogate node of i $Sg(i) := \operatorname{argmin} d_{ii}$

$$j \in V$$



If (i, j) is an edge in the latent tree, then (Sg(i), Sg(j)) is an edge in the Chow-Liu tree

CLGrouping Algorithm

RG ➡

RG

Step 1. Using information distances of observed nodes, construct MST(V; D). Identify the set of internal nodes.

Step 2. Select an internal node and its neighbors, and apply the recursive-grouping (RG) algorithm.

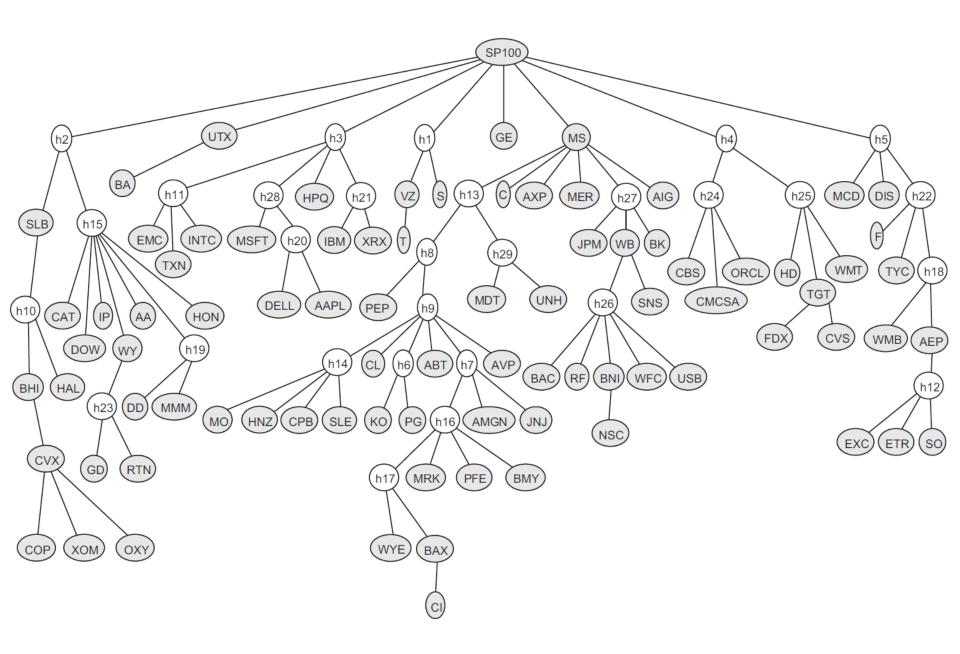
Step 3. Replace the output of RG with the sub-tree spanning the neighborhood.

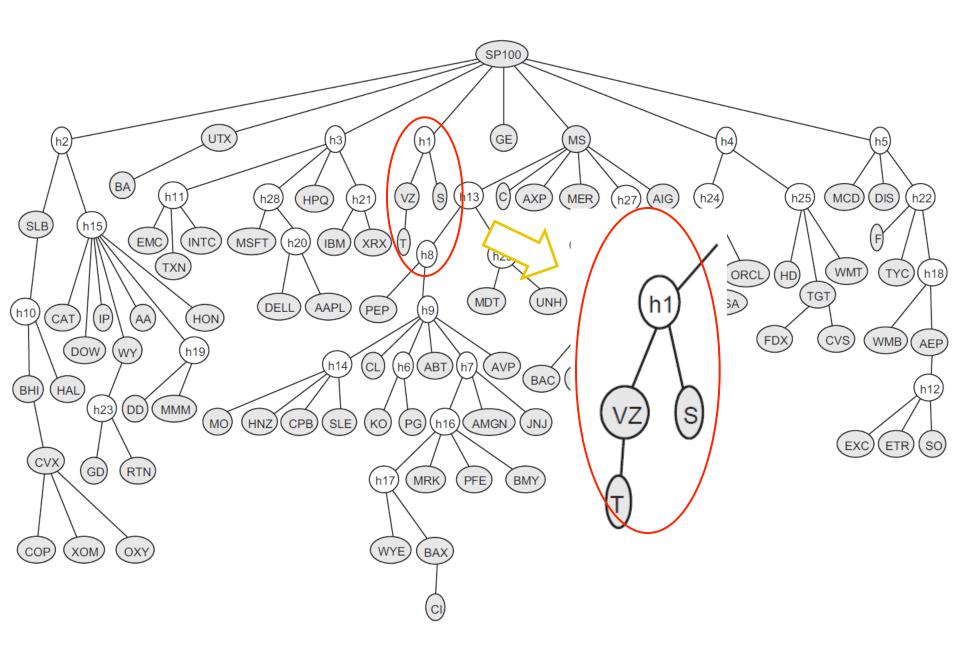
Repeat Steps 2-3 until all internal nodes are operated on. Computational complexity O(m² log m + (#internal nodes) (maximum degree)³)

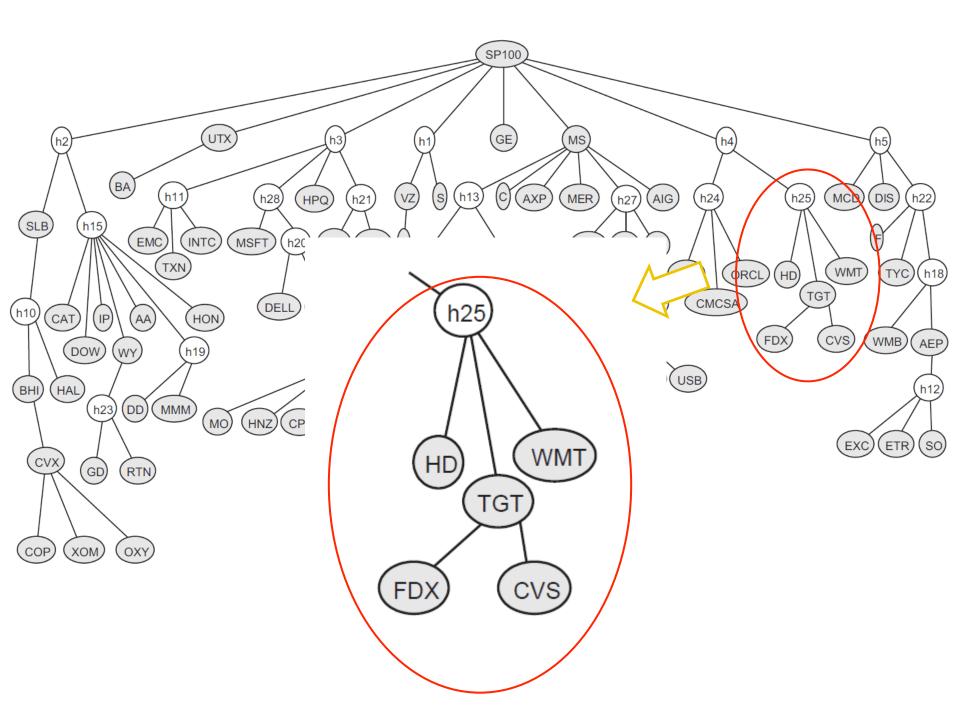
Sample-based Algorithms

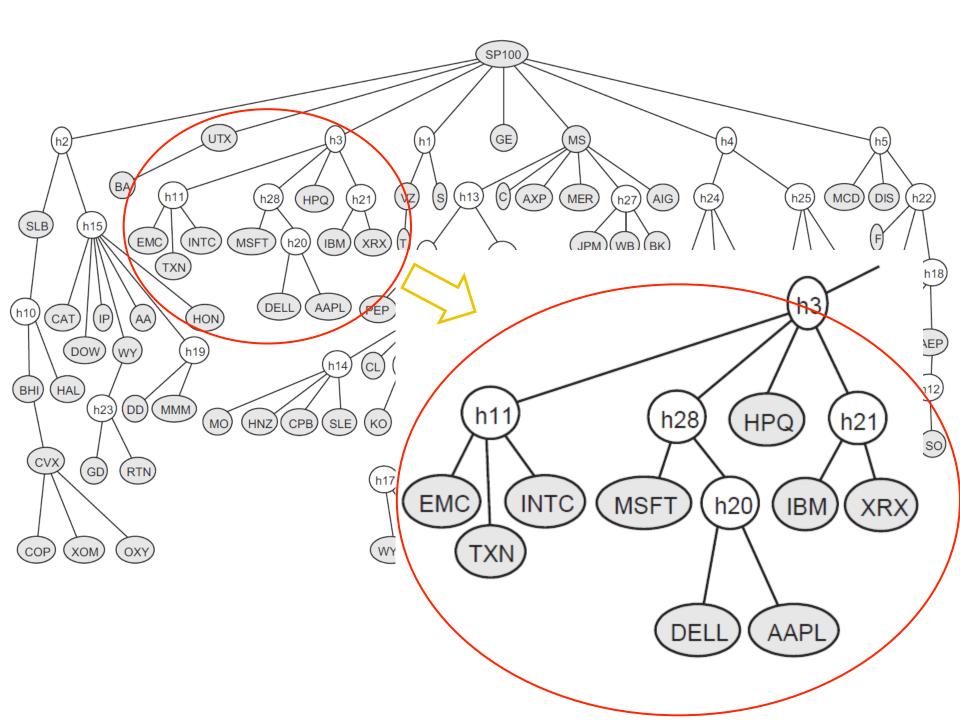
- Compute the ML estimates of information distances.
- Relaxed constraints for testing node relationships.
- Consistent (only in structure for discrete distributions)

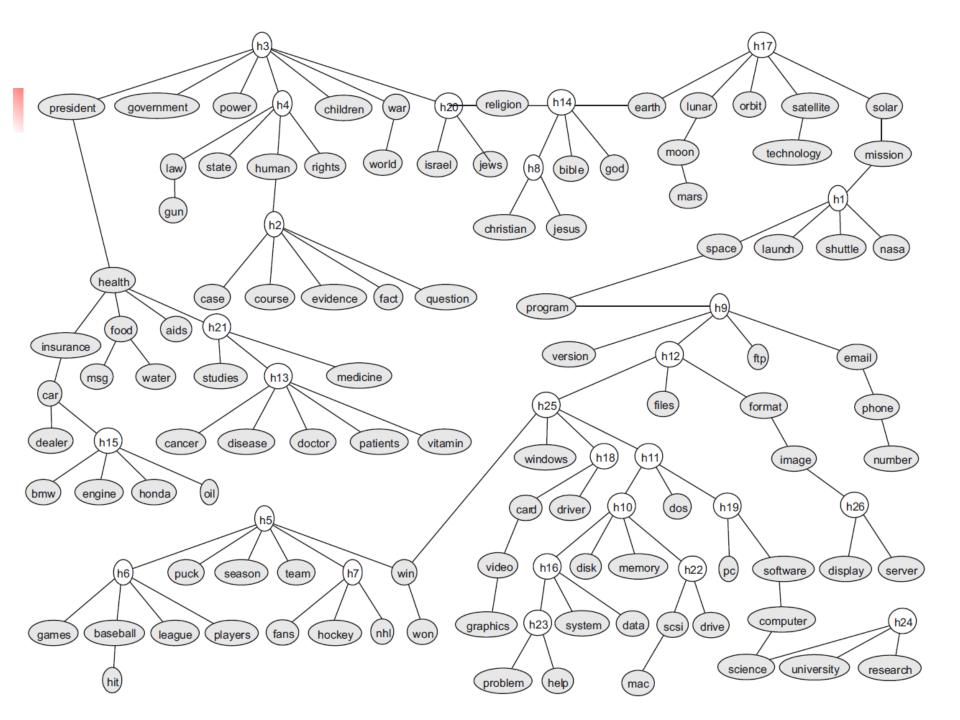
 Regularized CLGrouping for learning latent tree approximations.

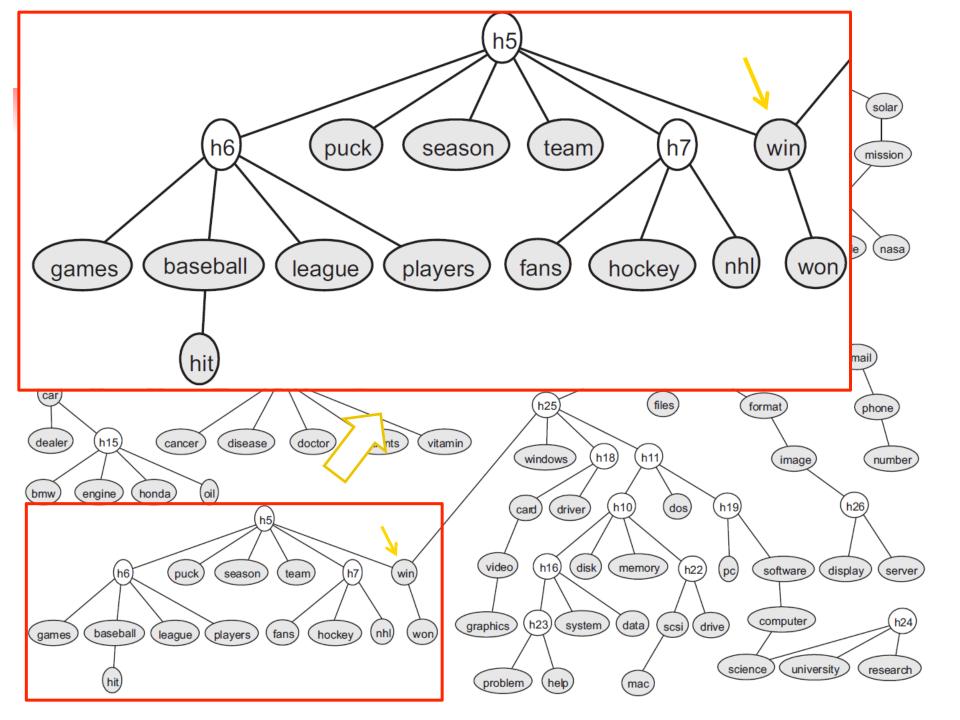


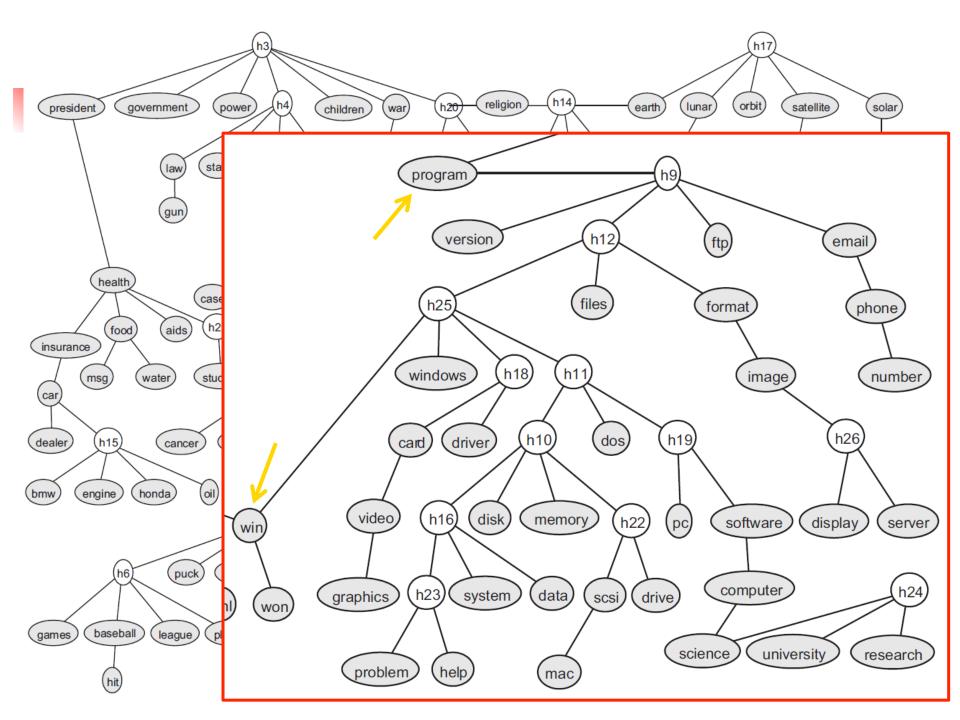


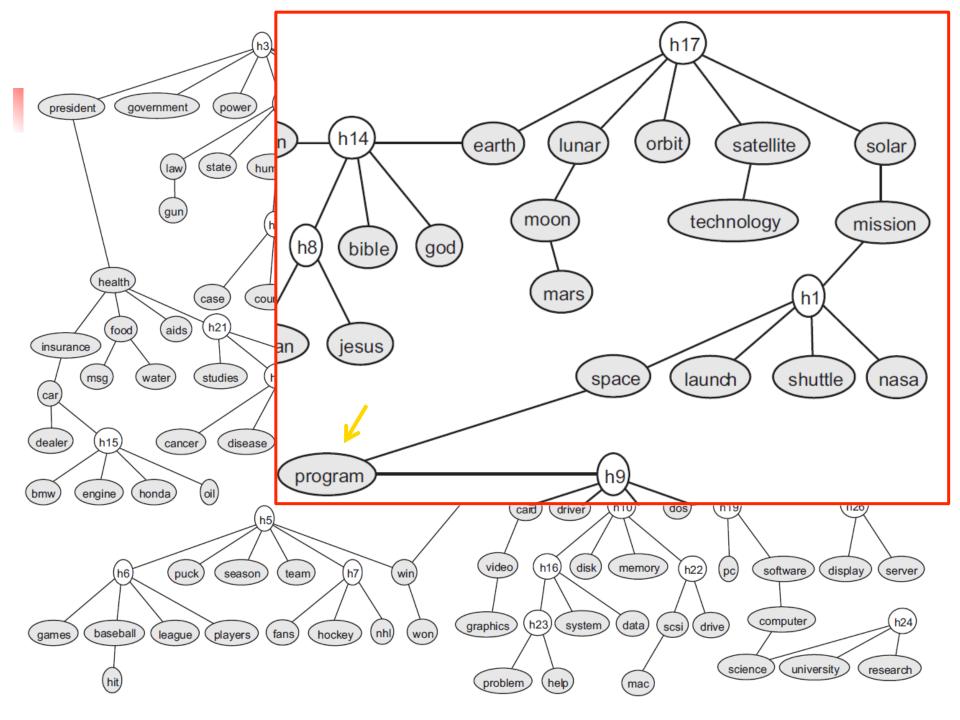


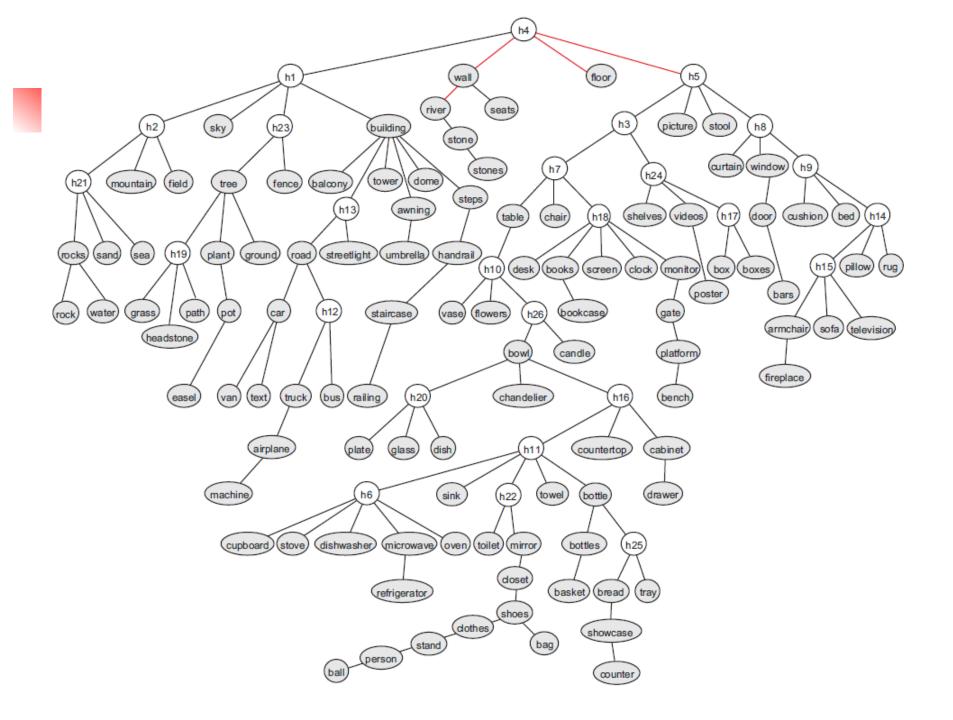


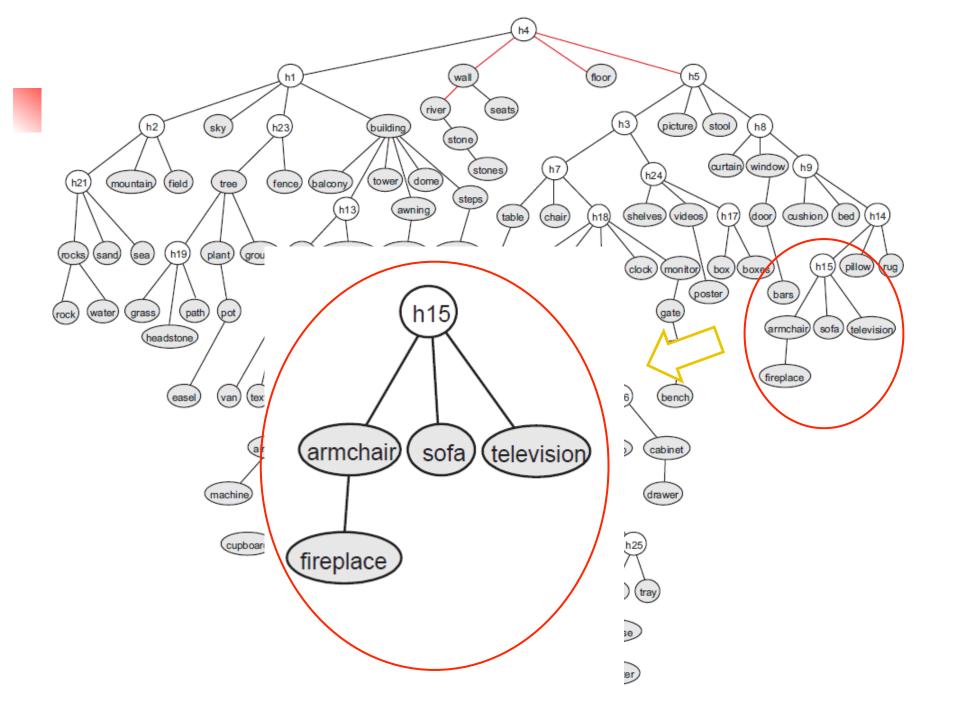


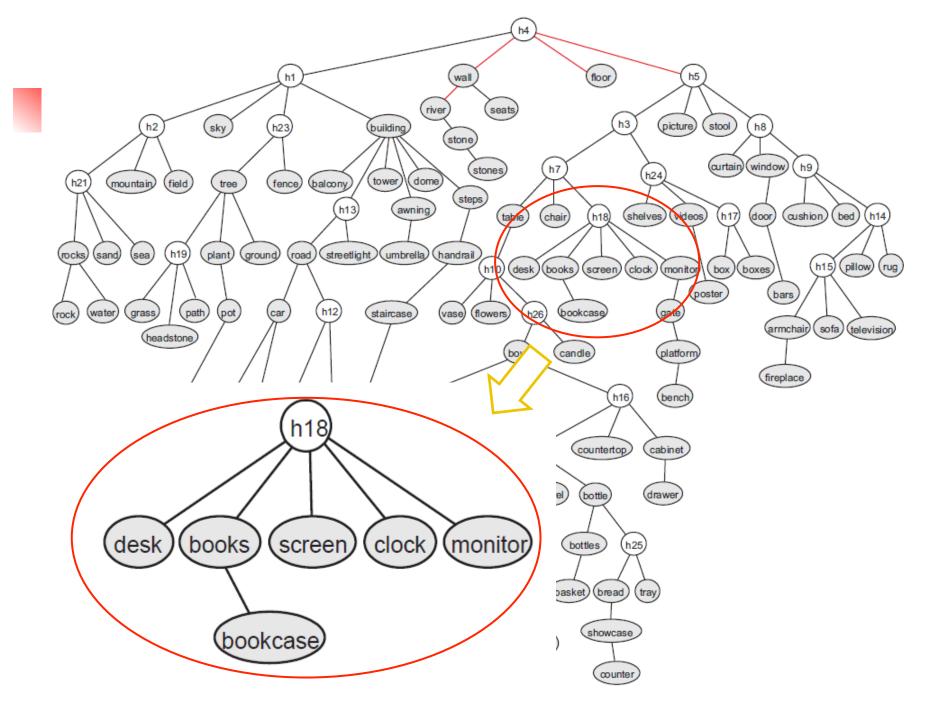






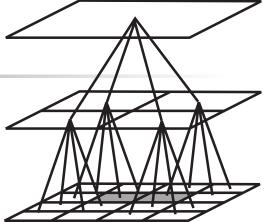






The dark side of trees = The bright side: No loops

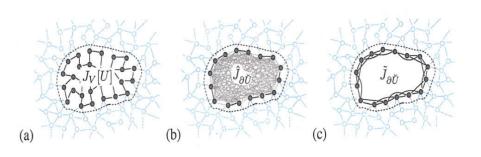
- So, what do we do?
- Try #1: Turn it into a junction tree
 - Not usually a good idea, but ...

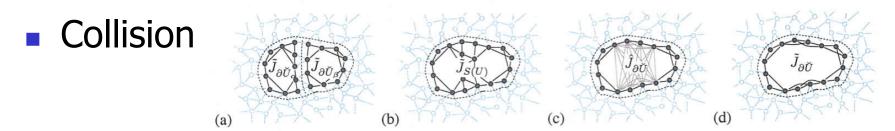


- Try #2: Pretend the problem isn't there and use a tree
 - If the real objectives are at coarse scales, then fine-scale artifacts may not matter
- Try #3: Pretend it's a tree and use (Loopy) BP
- Try #4: Think!
 - What does LBP do?
 - Better algorithms?
 - Other graphs for which inference is scalable?

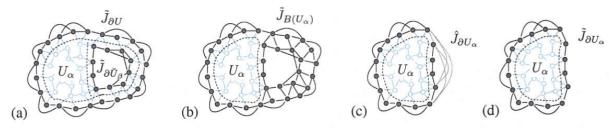
Recursive Cavity Models: "Reducedorder" modeling as part of estimation

Cavity thinning





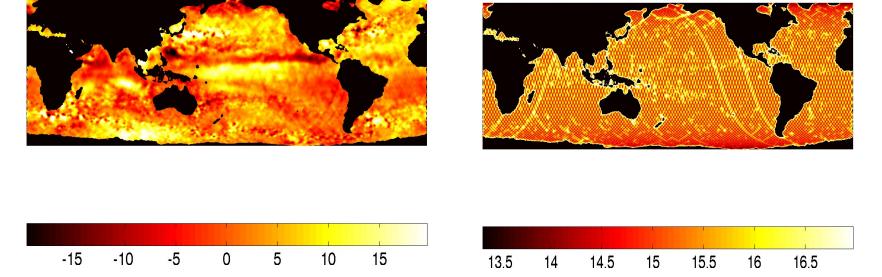
Reversing the process (bring your own blanket)



RCM in action: We are the world

SSHA Estimation Error (mm)

Estimated SSHA (cm above Mean-Sea-Level)



- This is the information-form of RTS, with a thinning approximation at each stage
- How do the thinning errors propagate? A control-theoretic stability question

Walk-sums and Gaussian models

$$J = P^{-1} \qquad \qquad h = P^{-1}\mu$$

Assume J normalized to have unit diagonal

$$J^{-1} = (I - R)^{-1} = I + R + R^2 + \dots$$

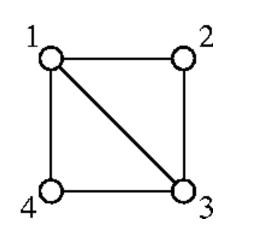
- *R* is the matrix of *partial correlation coefficients*
- $(R^{\ell})_{s,t} = sum over weighted length-\ell walks$ from s to t in graph

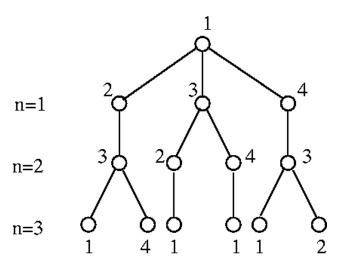
$$P_{s,t} = \phi(s \to t), \quad \mu_t = \sum_{s \in V} h_s \phi(s \to t)$$

- Inference algorithms may "collect" walks in different ways
- Walk-summability, corresponding to $\rho(\bar{R}) < 1$, guarantees
 - Collection in any order is OK
 - LBP converges
 - If LBP converges it collects all walks for μ_i but only **some** of the **self-return walks** required for P_{ii}

• There are lots of interesting/important models that are non-WS (and for which BP goes haywire)

A computation tree





- BP includes the *back-tracking self-return* walk (1,2,3,2,1)
- BP does *not* include the walk (1,2,3,1)
- BUT: For Non-WS models, the tree may be nonsensical
- There are ways to collect some or all of the missed walks
 - Embedded subgraphs as preconditioners
 - Convergence for WS models always

•A method that works also for non-WS models, recognizing that not all nodes are created equal

An alternate approach: Using (Pseudo-) *Feedback Vertex Sets*

- Provide additional potentials to allow computation of quantities needed in mean/variance/covariance computation in the FVS
- Run BP with both original potentials and the additional set(s)
- Feed back information to FVS to allow computation of exact variance and mean within the FVS
- Send modified information potentials to neighbors of FVS
- Run BP with modified information potentials
 - Yields exact means immediately
 - Combining with results from Steps 2, 3 yields exact variances

Approximate FVS

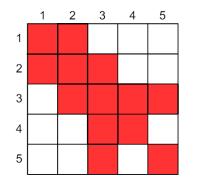
- Complexity is $\mathcal{O}(k^2n)$, where $k = |\mathcal{F}|$
- If k is too large
 - Use a pseudo- (i.e., partial) FVS, breaking only some loops
 - On the remaining graph, run LBP (or some other algorithm)
- Assuming convergence (which does *not* require WS)
 - Always get the correct means and variances on \mathcal{F} , exact means on \mathcal{T} , and (for LBP) approximate variances on \mathcal{T}
 - The approximate variances collect more walks than LBP on full graph
- Local (fast) method for choosing nodes for the pseudo-FVS to:
 - Enhance convergence
 - Collect the most important wants
- Theoretical and empirical evidence show $k \approx O(\log n)$ works

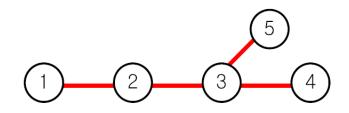
Motivation from PDEs: MultiPOLE Models

- Motivation from methods for efficient preconditioners for PDEs
 - Influence of variables at a distance are wellapproximated by coarser approximation
 - We then only need to do *LOCAL* smoothing and correction
- The idea for statistical models:
 - Pyramidal structure in scale
 - However, when conditioned on neighboring scales, the remaining correlation structure at a given scale is sparse and local

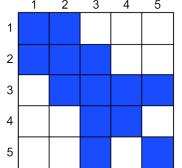
Models on graphs and on conjugate graphs

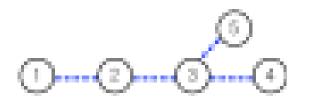
Garden variety graphical model: sparse *inverse* covariance

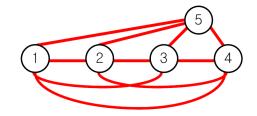




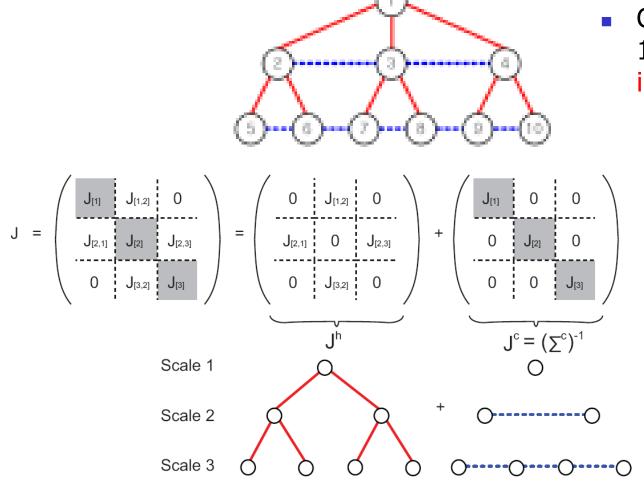
Conjugate models: sparse *covariance*







Inspiration from Multipole Methods for PDEs: Allow *Sparse Residual Correlation* Within Each Scale



 Conditioned on scale 1 and scale 3, x₂ is independent of x₄.

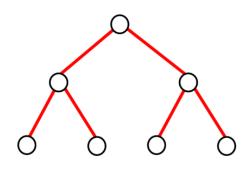
Learning such models:

"Dual" convex optimization problems

Multipole Estimation

• Richardson Iteration to solve $(J^h + (\Sigma^c)^{-1})x = h$

- Global tree-based inference
- Sparse matrix multiplication for in-scale correction



$$J^{h}x_{new} = h - (\Sigma^{c})^{-1}x_{old}$$

Compute last term via sparse equation

$$\Sigma^{c} Z = X_{old}$$

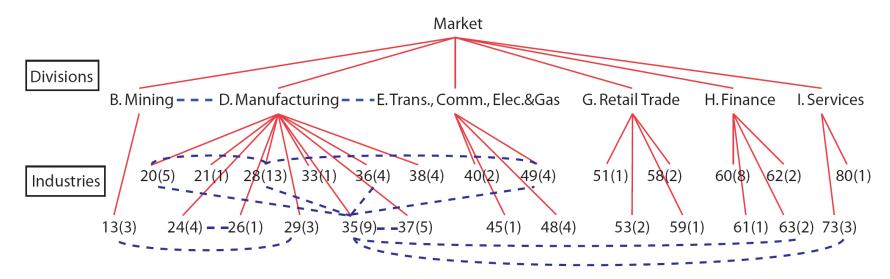


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 $X_{new} = \Sigma^{c}(h - J^{h}X_{old})$



Stock Returns Example



- Monthly returns of 84 companies in the S&P 100 index (1990-2007)
- Hierarchy based on the Standard Industrial Classification system
- Market, 6 divisions, 26 industries, and 84 individual companies
- Conjugate edges find strong residual correlations
 - Oil service companies (Schlumberger,...) and oil companies
 - Computer companies, Software companies, electrical equipment
 - . . .

What If Some Phenomena Hidden? United American Oil Chrysler United Chrysler JetBlue American GM GM Delta JetBlue Continental Ford Ford Delta

Continental

- Many dependencies among observed variables
- Less concise model

Hidden variables

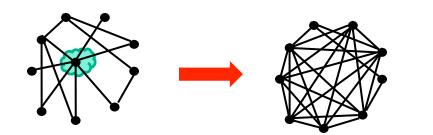
- Hedge fund investments,
- Patent accepted/rejected,
- Geopolitical factors,
- Regulatory issues, ...

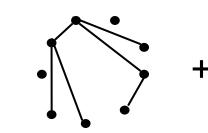
Graphical Models With Hidden Variables: Sparse Modeling Meets PCA

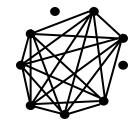
$$X = (X_O, X_H) \sim \mathcal{N}(0, \Sigma)$$

 $\Sigma = \begin{bmatrix} \Sigma_O & \Sigma_{O,H} \\ \Sigma_{H,O} & \Sigma_H \end{bmatrix} \qquad \qquad \Sigma^{-1} = K = \begin{bmatrix} K_O & K_{O,H} \\ K_{H,O} & K_H \end{bmatrix}$

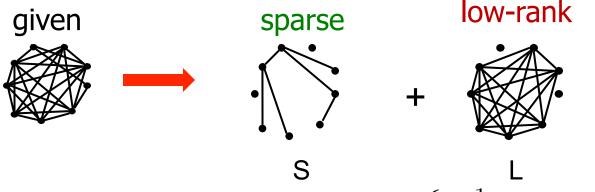
Marginal concentration $(\Sigma_O)^{-1} = K_O - K_{O,H} K_H^{-1} K_{H,O}$ matrix * * Sparse Low-rank







Convex Optimization for Modeling



• Samples of obs. vars.: $\mathcal{D}_n = \{X_O^1, \dots, X_O^n\}$

$$(\hat{S}_n, \hat{L}_n) = \arg\min_{S,L} -\frac{1}{n}\log \text{lik.} (S-L; \mathcal{D}_n) + \lambda_n [\gamma \|S\|_1 + \text{trace}(L)]$$

s.t. $S-L \succ 0, L \succeq 0.$

- Last two terms provide convex regularization for sparsity in *S* and low-rank in *L*
- Weights allow tradeoff between sparsity and rank

When does this work?

Identifiability conditions

- The sparse part can't be low rank and the low rank part can't be sparse
- There are precise conditions (including conditions on regularization weights) that guarantee
 - Exact recovery if exact statistics are given
 - Consistency results if samples are available ("with high probability" scaling laws on problem dimension and available sample size)

On the way: Construct richer classes of models for which inference is **easy**

- We have
 - Methods to learn hidden trees with fixed structure but unknown variable dimensions
 - Method to learn hidden trees with unknown structure but fixed (scalar) variable dimension
- Can we do both at the same time?
- We have method for recovering sparse plus low rank
 - How about tree plus low rank (i.e., small FVS)?
- Message-passing algorithms are distributed dynamic systems. How about designing better ones than LBP?